

Systematic review and meta-analysis of the prognostic role of fibroblast-activation protein in gastrointestinal cancers

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Background: Gastrointestinal (GI) cancers, particularly pancreatic cancer, are characterized by a dense stromal tumor microenvironment where cancer-associated fibroblasts (CAFs) predominate. CAFs comprise highly heterogeneous subpopulations with different functions, which can be both tumor-promoting and tumor-restraining. This systematic review and meta-analysis aims to comprehensively assess the impact of the CAF marker fibroblast-activation protein (FAP) expression on clinical outcomes in GI cancers.

Methods: Adhering to PRISMA guidelines, we systematically searched PubMed/MEDLINE, Web of Science, Cochrane Library, and ClinicalTrials.gov for relevant articles. Inclusion criteria involved studies comparing GI cancer patients with and without FAP overexpression. Meta-analysis evaluated overall survival (OS), histological differentiation, local tumor invasion, lymph node metastases, and distant metastases. For each observational study, the risk of bias was assessed using the risk of bias in non-randomized studies of exposure (ROBINS-E) tool.

Results: The meta-analysis included 31 cohort studies from six countries, comprising 3,976 patients. Patients without FAP overexpression exhibited a favorable OS [hazard ratio (HR) =1.74; 95% confidence interval (CI): 1.51–2.01]. Subgroup analyses revealed consistent results across esophageal, pancreatic, colorectal, and gastric cancers. While one-year survival rates showed no significant difference, subsequent years displayed lower rates for FAP-overexpressing groups. Lymph node metastases were more frequent in FAP-overexpressing patients, whereas distant metastases did not differ. None of 31 studies systematically controlled confounding and adjusted data so that all studies were categorized as "high risk of bias" for the domain "risk of bias due to confounding". For domains "risk of bias arising from measurement of exposure", "risk of bias due to post-exposure interventions", "risk of bias arising from measurement of outcomes", and "risk of bias in selection of the reported result", all studies were categorized as "low risk of bias".

Conclusions: This meta-analysis underscores the potential adverse prognostic significance of FAP expression in GI cancers. Limitations include heterogeneity in FAP expression cutoffs and definitions. Future research should focus on delineating the precise roles and clinical implications of FAP in GI cancers.

Keywords: Fibroblast-activation protein (FAP); cancer-associated fibroblasts (CAFs); gastrointestinal cancers (GI cancers); meta-analysis

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Introduction

Cancers of the gastrointestinal (GI) tract include esophageal, pancreatic, colorectal, gastric, and liver cancer. These five major types of GI cancer represent 26% of the global cancer incidence and 35% of all cancer related deaths in 2018 (1). A hallmark of GI cancers, especially pancreatic cancer, is the dense and complex stromal tumor microenvironment where cancer-associated fibroblasts (CAFs) are the predominant stromal cell type (2). They comprise highly heterogeneous subpopulations with diverse and sometimes opposing functions, ranging from tumorpromoting to tumor-restraining roles (3). CAFs contribute to the reprogramming of the immune microenvironment, thereby facilitating and promoting cancer proliferation, migration, invasion, as well as metastasis (4,5). Given their influence on tumor progression, CAFs have been proposed as therapeutic target in GI cancers, especially in pancreatic cancer. However, the depletion of α -smooth muscle actin [(α-SMA), a major CAF marker]-positive cells in pre-

Highlight box

Key findings

 Fibroblast-activation protein (FAP) overexpression in gastrointestinal (GI) cancers is associated with poorer overall survival (OS), as shown in 31 cohort studies involving 3,976 patients. This association was consistent across several GI cancers, including esophageal, pancreatic, colorectal, and gastric cancers. Furthermore, patients with FAP overexpression exhibited a higher frequency of lymph node metastases compared to those without FAP overexpression.

What is known and what is new?

- FAP is a marker predominantly observed in cancer-associated fibroblasts (CAFs) within the tumor microenvironment of GI cancers. Its expression is linked to adverse outcomes, including poorer OS and increased lymph node metastases.
- The findings highlight the need for standardized definitions and cutoff values for FAP expression in research.

What is the implication, and what should change now?

• The study underscores the potential of FAP as a prognostic marker and a therapeutic target in GI cancers. The variability in FAP expression and its impact on survival across different GI cancers calls for further investigation into its role and the development of selective FAP inhibitors or targeted therapies. Future research should focus on identifying and targeting specific FAP-positive CAF subtypes that contribute to tumor progression. Standardization in measuring FAP expression and controlling for confounding factors in studies are critical steps to improve the reliability of research in this area. clinical models of pancreatic cancer has paradoxically led to invasive, undifferentiated tumors and reduces animal survival, highlighting the complexity of targeting CAFs due to their functional heterogeneity (6). As a result, identification and specific targeting of tumor-promoting CAF subtypes and markers are emerging strategies (7). To identify and characterize CAF subtypes, a number of markers have been identified, such as desmin, fibroblastactivation protein (FAP), fibroblast-specific protein (FSP), podoplanin (PDPN), α -SMA, and vimentin (8,9).

FAP, a type II trans-membrane serine protease, which shares high sequential similarity with the dipeptidyl peptidase (DPP) 4, is particular interest (10). Although expression of FAP is typically low to undetectable in most adult tissues, it becomes markedly upregulated in multiple types of cancer, and it is predominantly observed in CAFs (11). In a pre-clinical study, it has been shown that FAP activates macrophages and exacerbates liver inflammation and fibrosis (12). FAP plays a key role in promoting tumor progression and metastasis, further shaping the immunosuppressive tumor microenvironment (13). Inhibition or deletion of FAP-positive CAFs in pancreatic cancer models has been associated with reduced tumor growth and improved survival (14,15). Recent findings further differentiate the roles of tumor-promoting FAPpositive CAFs from tumor-restraining α-SMA-positive CAFs, emphasizing the need for precise targeting strategies (15).

To assess the currently available evidence on the impact of FAP expression on survival and clinical characteristics in GI cancers, we performed a systematic review and meta-analysis. We present this article in accordance with the PRISMA reporting checklist (available at https://jgo. amegroups.com/article/view/10.21037/jgo-24-374/rc) (16).

Methods

The study has been registered in the PROSPERO (International Prospective Register of Systematic Reviews) database (CRD42022372194) (17). The protocol of this meta-analysis was published a priori (18).

The inclusion criteria were defined according to the Population, Intervention, Comparison, Outcomes and Study (PICOS) principles. The population consisted of adults with GI cancers exhibiting FAP overexpression. The intervention was defined as diagnostic methods utilizing FAP. The comparator were patients without FAP overexpression. The outcome measures included postoperative survival, histological differentiation, local tumor invasion, lymph



Figure 1 PRISMA flow chart.

node metastases, and distant metastases. The included study designs were randomized controlled trials (RCTs), cohort studies, case-control studies, and cross-sectional studies relevant to FAP in GI cancers.

Search strategy

The databases PubMed/MEDLINE, Web of Science Core Collection, Cochrane Library, and ClinicalTrials.gov were searched via their respective online search engines. Citavi 6 (Swiss Academic Software GmbH) was used as an automatic deduplication system for the studies retrieved from the several databases (*Figure 1*). The search was performed on studies published until December 29, 2022. The search strategies used in each database are displayed in Appendix 1. Titles and abstracts were evaluated independently in a standardized manner by two authors (A.R. and Y.S.) to assess eligibility for inclusion. All the potential studies identified from the search were coded as either "retrieve" (eligible, potentially eligible, or unclear) or "do not retrieve". For studies coded "retrieve", two reviewers (A.R. and Y.S.) independently screened the full text and recommended inclusion or exclusion. Disagreements between the reviewers was resolved by consensus; if no agreement was reached, a third reviewer (R.B.) decided whether to include the respective study. The reference lists of the included studies were manually searched to find additional relevant articles.

Inclusion and exclusion criteria

Only articles in English were considered. Studies comparing patients with and without FAP overexpression (regardless of the specific cutoff value used) in GI tumors and reporting on at least one of the following a priori defined outcomes were included: postoperative survival (overall and median survival, 1-, 2-, 3-, and 5-year survival rates), histological differentiation (grading), local tumor invasion (as defined in the included studies), lymph node metastases, and distant metastases. Review articles, case reports, case series with less than five patients, commentaries, and letters were not included. Details of the study selection process were summarized in a flowchart according to the recommendations of the PRISMA 2020 statement.

Data collection

Data from the included studies were extracted separately by two authors (Y.S. and A.R.) and stored in a dedicated database. The following descriptive data were documented for each selected study: first author, year of publication, inclusion period, country/region and city where the study was conducted, sample size, and mean or median follow-up time. The distribution of the following patient characteristics was documented: tumor type, histopathological tumor stage [using the Union for International Cancer Control (UICC) tumor, node, metastasis, grade (TNMG) classification system], presence and type of neoadjuvant therapy, presence and type of adjuvant therapy, FAP detection method, FAP antibody, FAP location, number of FAP-positive cases, and cutoff for overexpression. The following predefined outcomes were extracted: postoperative survival (overall and median survival, 1-, 2-, 3-, and 5-year survival rates), histological differentiation (grading), local tumor invasion (as defined in the included studies), lymph node metastases, and distant metastases. Subgroup analysis was performed for location of FAP expression (tumor stroma or tumor cells or both) and tumor type (esophageal cancer, pancreatic cancer, gastric cancer, colorectal cancer, hepatocellular carcinoma, and cholangiocellular carcinoma).

For each observational study, the risk of bias was assessed using the risk of bias in non-randomized studies of exposure (ROBINS-E) tool (19).

Statistical analysis

A meta-analysis comparing patients with and without FAP overexpression with the following outcomes was conducted: postoperative survival (overall and median survival, 1-, 2-, 3-, and 5-year survival rates), histological differentiation (grading), local tumor invasion (as defined in the included studies), lymph node metastases, and distant metastases. Time-to-event outcomes were analyzed using hazard ratios (HRs) with 95% confidence intervals (CIs). For estimating pooled overall effects, we used both random- and common-effect models. The Review Manager (RevMan) software (version 5.4; Cochrane Collaboration) was used. The

magnitude of the effect estimate was visualized by forest plots. Odds ratios (ORs) were calculated for binary data, and weighted mean differences and relative standard deviation were determined for continuous data. For calculating ORs, random-effect models were used. The 95% CI, heterogeneity, and statistical significance were reported for each outcome. The Chi-squared test and degrees of freedom (df) were used to evaluate heterogeneity and statistical significance. A P value of <0.05 was considered statistically significant. Heterogeneity was presented as I^2 . I^2 >75% was defined as "considerable heterogeneity" (https:// training.cochrane.org/handbook/current/chapter-10). If the study reported overall survival (OS) in the form of a Kaplan-Meier estimate and no specific data were available, then the graphically presented data were extracted using the Enguage Digitizer 12.1 (https://github.com/markummitchell/ engauge-digitizer), and the HR and standard error (SE) was calculated using a spreadsheet designed by Tierney (20). For sensitivity analyses, all studies with a high or serious risk of bias were excluded, and the analyses of the outcomes, as described above, were conducted.

Results

Thirty-one cohort studies [4 studies on esophageal cancer (21-24), 8 studies on pancreatic cancer (25-32), 9 studies on colorectal cancer (33-41), 7 studies on gastric cancer (42-48), 2 studies on hepatocellular carcinoma (49,50), and 1 study on cholangiocarcinoma (51)] from 6 countries published between 2007 and 2022, were included in the meta-analysis (*Figure 1*). The enrolment periods of these studies ranged from 1981 to 2018. In these studies, a total of 3,976 patients (2,114 with FAP overexpression and 1,862 without overexpression) were included. The study features, patient characteristics, follow-up and outcomes are presented in table available at https://cdn.amegroups.cn/static/public/jgo-24-374-1.xlsx.

No subgroup analysis was conducted regarding the location of FAP expression due to insufficient information provided. Additionally, a meta-analysis could not be performed for histological differentiation (grading) and local tumor invasion due to inadequate information or heterogeneous definitions provided. Regarding OS, patients with FAP overexpression experienced a poorer outcome than those without overexpression (HR =1.74; 95% CI: 1.51–2.01; heterogeneity χ^2_{30} =79.94; P<0.001; I²=62%) (*Figure 2*). This result persisted within specific cancer subgroups: esophageal cancer (HR =2.06; 95%)



Figure 2 Forest plot of HR with 95% CI for the 31 studies as well as for pooled HR for all studies calculated by using the common and

random effect models. The HRs presented are FAP overexpression *vs.* no FAP overexpression. Heterogeneity was presented as I². HR, hazard ratio; CI, confidence interval; FAP, fibroblast-activation protein.

CI: 1.53–2.77; heterogeneity χ_{3}^{2} =6.41; P=0.09; I²=53%), pancreatic cancer (HR =2.06; 95% CI: 1.45–2.92; heterogeneity χ_{7}^{2} =12.76; P=0.08; I²=45%), colorectal cancer (HR =1.53; 95% CI: 1.18–1.97; heterogeneity χ_{8}^{2} =17.25; P=0.03; I²=54%), and gastric cancer (HR =1.85; 95% CI: 1.48–2.32; heterogeneity χ_{6}^{2} =2.4; P=0.88; I²=0%). For the subgroup hepatocellular carcinoma there was no significant association between FAP overexpression and survival. In addition, the data was highly heterogeneous (HR =1.32; 95% CI: 0.33–5.22; heterogeneity χ_{1}^{2} =7.01; P=0.008; I²=86%). We did not conduct a sensitivity analysis for the subgroup hepatocellular carcinoma, since we identified only two studies. Notably, only the study conducted by Byrling in 2020 investigated cholangiocarcinoma, reporting that patients without FAP overexpression exhibited shorter survival (HR =0.74; 95% CI: 0.40-1.36) (*Figure 3*). The range of median overall survival (mOS) for patients with overexpression and without overexpression was 10.0 to 125.9 months and 16.6 to 81.5 months, respectively. Analyzing the available dichotomous survival data, identical 1-year survival rates were noted in patients with esophageal cancer between the overexpression and no overexpression groups (data from 4 studies, 78% vs. 80%, OR =0.93; 95% CI: 0.56-1.53; P=0.76). In the subgroup of pancreatic



Figure 3 Forest plot of HR with 95% CI for the 31 studies as well as for GI cancer subgroup-dependent pooled HR calculated by using the common and random effect models. The HRs presented are FAP overexpression *vs.* no FAP overexpression. Heterogeneity was presented as I². HR, hazard ratio; CI, confidence interval; GI, gastrointestinal; FAP, fibroblast-activation protein.

cancer, lower 1-year survival rates were evident in the overexpression group (data from five studies, 59% *vs.* 73%, OR =0.66; 95% CI: 0.45–0.97; P=0.03). Similarly, lower 2- and 3-year survival rates were noted in the overexpression group for both esophageal and pancreatic cancer (41% *vs.* 60%, OR =0.62, 95% CI: 0.47–0.81, P<0.05; 25% *vs.* 57%, OR =0.25, 95% CI: 0.18–0.35, P<0.05). It is noteworthy that 5-year survival rates were higher in the FAP overexpression group for both subgroups (39% *vs.* 34%, OR =1.71; 95% CI: 1.26–2.32, P<0.05) (Figures S1-S4).

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Regarding lymph node metastases, a higher frequency was evident in the FAP overexpression group compared to the non-overexpression group (56% vs. 44%, OR =1.98; 95% CI: 1.59–2.46; P<0.05) (*Figure 4*). In terms of distant metastases, there was no difference between the two groups (11% vs. 12%, OR =0.92; 95% CI: 0.51–1.64; P=0.77).

ROBINS-E bias analysis revealed that none of 31 studies systematically controlled confounding and adjusted data via e.g., propensity score matching, so that all studies were categorized as "high risk of bias" for the domain "risk of bias due to confounding" (Figure 5). Currently, there is no evidence identifying potential confounding factors for FAP expression. As all 31 studies are observational and not RCTs, acquiring comprehensive data on confounding, including all unknown factors, is unattainable. Yet, several studies stratified FAP expression status and analyzed several potential confounders such as age or gender, and no study detected a notable difference in age or gender between FAP overexpression and no overexpression groups (21-24,26-28,31,32,34,35,38,41-43,45-47,49-51). These studies were categorized as "low risk of bias" for the domain "risk of bias due to missing data", and the other studies were categorized as "some concerns". All 31 studies were categorized "some concerns" for the domain "risk of bias in selection of participants into the study". General selection bias is discussed in the discussion section. For domains "risk of bias arising from measurement of exposure", "risk of bias due to post-exposure interventions", "risk of bias arising from measurement of outcomes", and "risk of bias in selection of the reported result", all studies were categorized as "low risk of bias" (Figure 5).

Discussion

Increasing evidence suggests that the tumor microenvironment and CAFs, a pivotal component of the tumor stroma, play critical roles in facilitating cancer progression, metastasis, drug resistance, and immunosuppression in many types of cancer including GI cancer (52). CAFs can originate from diverse types of cells such as adipocytes and mesenchymal stem cells (3). CAFs are a highly heterogeneous group of cells and exhibit high cellular plasticity (3). Therefore, identification of tumor-promoting CAF subtypes and specific markers are emerging to develop precise targeting strategies. FAP has been considered to label tumorpromoting CAF subtypes and to play a key role in promoting tumor progression and metastasis, further shaping the immunosuppressive tumor microenvironment (13). In the current study, we conducted a systematic review and meta-analysis concerning FAP overexpression in GI cancers and observed that patients without FAP overexpression experienced a favorable outcome regarding OS compared to patients with FAP overexpression. Several studies on this topic have been published and the last meta-analysis published in 2015 included 15 studies with various cancer types. Eight studies involving 1,277 patients with GI cancers were included in this meta-analysis (2 esophageal, 1 gastric, 2 pancreatic, and 3 colorectal cancer). It reported an association between FAP overexpression and shorter OS in pancreatic cancer (HR =3.18; 95% CI: 1.42-7.12) and colorectal cancer (HR =1.72; 95% CI: 1.14-2.60) (53). In our meta-analysis, we identified 8 pancreatic and 9 colorectal cancer cohort studies, for which we demonstrated an association between FAP overexpression and shorter OS. We further observed that patients with esophageal cancer and gastric cancer who showed FAP overexpression experienced a shorter survival compared to those without overexpression. We identified only one study, in patients with cholangiocarcinoma, which yielded inconclusive results.

FAP has been considered one of the most relevant CAF markers. The question of whether FAP can be a viable target in the treatment of GI cancers is still not fully understood and requires further investigation. A pre-clinical study showed that stromal FAP promotes intrahepatic cholangiocarcinoma growth via chemokine CCL2 expression and recruitment of monocyte-derived suppressor cells (MDSCs) (54). Consistently, the chemokine-mediated immunosuppressive and tumor-promoting role of FAP-positive CAFs has been demonstrated in pancreatic cancer and liver cancer (55,56). These data suggest a tumor-promoting role of FAP-positive CAFs by expressing chemokines and establishing an immunosuppressive tumor microenvironment. On the other hand, several studies have shown the organ-specific variations of transcriptional

	FAP overexpre	ession	No FAP overexpr	ession		Odds Ratio	Odds Ratio		
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% Cl	M–H, Fixed, 95% Cl		
1.1.1 Esophageal car	ncer								
Higashino 2019	19	31	8	39	2.3%	6.14 [2.12, 17.73]			
Kashima 2019	44	50	15	44	1.6%	14.18 [4.93, 40.78]			
Li 2020	32	45	38	76	6.9%	2.46 [1.12, 5.40]			
Subtotal (95% CI)		126		159	10.9%	5.00 [2.95, 8.48]	•		
Total events	95		61						
Heterogeneity: Chi ² = Test for overall effect	7.00, df = 2 (P : Z = 5.97 (P < 0	= 0.03); 0.00001)	$l^2 = 71\%$						
1.1.2 Pancreatic cand	er								
Kawase 2015	17	29	9	19	3.8%	1.57 [0.49, 5.05]			
Mena 2022	37	51	58	91	9.7%	1.50 [0.71, 3.18]			
Shi 2012	35	98	10	31	8.3%	1.17 [0.49, 2.75]			
Wen 2019	22	33		34	2.2%	6.50 [2.22, 19.01]			
Subtotal (95% CI)		211		175	24.0%	1.86 [1.18, 2.93]	•		
Total events	111		85				-		
Heterogeneity: Chi ² = Test for overall effect	6.74, df = 3 (P : Z = 2.70 (P = 0	= 0.08); 0.007)	$I^2 = 56\%$						
1.1.3 Colorectal cand	er								
Hayase 2022	30	96	9	63	6.3%	2.73 [1.19, 6.24]			
Kim 2022	25	53	21	68	8.2%	2.00 [0.95, 4.21]			
Subtotal (95% CI)		149		131	14.5%	2.32 [1.33, 4.02]	•		
Total events	55		30						
Heterogeneity: $Chi^2 =$	0.30. df = 1 (P)	= 0.58):	$l^2 = 0\%$						
Test for overall effect	Z = 2.98 (P = 0)	0.003)							
1.1.4 Gastric cancer									
Fukumoto 2009	61	64	30	36	1.5%	4.07 [0.95, 17.39]	-		
Song 2017	20	70	22	42	16.6%	0.36 [0.16, 0.81]			
Sun 2022	35	42	37	51	4.7%	1.89 [0.68, 5.24]			
Tong 2022	54	68	54	103	7.5%	3.50 [1.73, 7.07]			
Subtotal (95% CI)		244		232	30.3%	1.56 [1.03, 2.37]	\bullet		
Total events	170		143						
Heterogeneity: Chi ² =	19.69, df = 3 (F	P = 0.000	$(12); 1^2 = 85\%$						
Test for overall effect	: Z = 2.09 (P = 0	0.04)							
1.1.5 Hepatocellular	carcinoma								
Miao 2014	22	36	31	50	8.5%	0.96 [0.40, 2.32]			
Zou 2018	8	64	7	74	4.8%	1.37 [0.47, 4.00]			
Subtotal (95% CI)		100		124	13.3%	1.11 [0.56, 2.19]	-		
Total events Heterogeneity: Chi ² = Test for overall effect	30 0.24, df = 1 (P : Z = 0.30 (P = 0	= 0.62); 0.77)	38 $I^2 = 0\%$						
1.1.6 Cholangiocarci	noma								
Byrling 2020	21	34	18	23	6.9%	0.45 [0.13, 1.50]			
Subtotal (95% CI)		34		23	6.9%	0.45 [0.13, 1.50]			
Total events	21		18				-		
Heterogeneity: Not an	plicable		-						
Test for overall effect	: Z = 1.30 (P = 0).19)							
Total (95% CI)		864		844	100.0%	1.98 [1.59, 2.46]	•		
Total events	482		375			- ,	•		
Heterogeneity: $Chi^2 =$	55.06. df = 15	(P < 0.00)	(001) : $I^2 = 73\%$						
Test for overall effect	7 = 6.15 (P < 0)	00001					0.01 0.1 1 10 100		
Test for solution of the form $(FAP +]$ Favours [FAP +] Favo									

Figure 4 Forest plot of OR with 95% CI for lymph node metastases. The ORs presented are FAP overexpression *vs.* no FAP overexpression. Heterogeneity was presented as I². FAP, fibroblast-activation protein; M-H, Mantel-Haenszel; CI, confidence interval; OR, odds ratio.

fingerprints in CAFs, especially for hepatic and pancreatic fibroblasts/CAFs (57,58). Therefore, we cannot exclude the possibility that FAP-positive CAFs in addition co-express organ-specific CAF markers, leading to functional diversity and disparity of FAP-positive CAFs between different GI tumor entities. To clarify this issue, it is necessary to conduct single nuclear RNA sequencing and compare gene expression profiling in CAFs from different GI tumor entities in the future.

The present study has some limitations. First, the cutoff values to define FAP overexpression and the antibodies employed for immunohistochemistry were heterogeneous between the single studies. This can increase the variation in results of the individual studies and make their

		DI	DO	K	ISK OT DIA	s domain	IS	D7	0 "
			D2	D3	D4	D5	D6		Overall
	Byrling et al. (2020)		(+)		(+)	+	$\mathbf{+}$	+	
	Chen <i>et al.</i> (2017)		+		(+)		(+)	(+)	X
	Cohen <i>et al.</i> (2008)	X	+	-	+	-	+	+	×
	Coto-Llerena et al. (2020)	X	+	-	+	+	+	+	×
	Fukumoto et al. (2009)	X	+	-	+	+	+	+	X
	Ha et al. (2014)	X	+	-	+	+	+	+	×
	Hayase et al. (2022)	X	+	-	+	+	+	+	X
	Henry <i>et al.</i> (2007)	X	+	-	+	-	+	+	X
	Herrera et al. (2012)	X	+	-	+	-	+	+	×
	Higashino <i>et al.</i> (2019)	X	+	-	+	+	+	+	×
	Kashima <i>et al.</i> (2019)	X	+	-	+	+	+	+	X
	Kawase et al. (2022)	X	+	-	+	+	+	+	X
	Kim <i>et al.</i> (2022)	X	+	-	+	+	+	+	X
	Kim et al. (2014)	X	+	-	+	+	+	+	X
	Li et al. (2020)	X	+	-	+	+	+	+	X
tudy	Lo <i>et al.</i> (2017)	X	+	-	+	+	+	+	X
Ó	Ma et al. (2018)	X	+	-	+	+	+	+	X
	Meng et al. (2022)	X	+	-	+	+	+	+	X
	Miao et al. (2014)	X	+	-	(+)	-	(+)	+	X
	Ogawa et al. (2021)	X	+	<u> </u>	+	-	+	+	X
	Park et al. (2017)	X	+	-	(+)	-	+	+	X
	Shi et al. (2021)	X	+	-	+	+	+	+	X
	Solano-iturri <i>et al.</i> (2020)	X	+	-	+	-	+	+	X
	Son et al. (2019)	X	+	-	+	-	+	(+)	X
	Song et al. (2017)	X	+	-	(+)	+	(+)	(+)	X
	Sun et al. (2022)	X	(+)	-	(+)	(+)	(+)	(+)	X
	Tong et al. (2022)	X	$\overline{+}$	-	(+)	(+)	(+)	(+)	X
	Wen et al. (2017)	X	+	-	(+)	-	(+)	(+)	X
	Wen et al. (2019)		$\overline{+}$		(+)	+	(+)	(+)	
			$\overline{+}$		$\overline{+}$	$\overline{+}$	$\overline{+}$		
			$\overline{+}$		$\overline{+}$	$\overline{+}$			
	Zou et al. (2018)	Domains:						ludger	ont
		D1: Bias du	ue to confo	unding.	at of the				ah
		D3: Bias in	selection o	f participan	ts into the s	study. (or int	o the analys	sis). 🦲 So	me concerns
		D4: Bias di D5: Bias di	ue to post-e ue to missin	exposure int ng data.	erventions.			+ Lo	w
		D6: Bias ar D7: Bias in	rising from r selection o	measurement of the report	nt of the ou ed result.	tcome.			

Figure 5 Summary of ROBINS-E bias analysis for the 31 studies. Risks were rated as "low risk of bias", "some concerns", or "high risk of bias". D1: risk of bias due to confounding. D2: risk of bias arising from measurement of the exposure. D3: risk of bias in selection of participant into the study. D4: risk of bias due to post exposure interventions. D5: risk of bias due to missing data. D6: arising from measurement of outcomes. D7: risk of bias risk of bias in selection of the reported result. D, domain; ROBINS-E, risk of bias in non-randomized studies of exposure.

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interpretation difficult. Second, the studies recruited only GI cancer patients and no control cohort. The inclusion of all GI cancer subtypes might have caused clinical heterogeneity, which we aimed to reduce by also conducting meta-analyses in subgroups of tumor subtype. Third, all included studies are observational and retrospective, and selection bias is substantial. The research has primarily focused on individuals who underwent surgical procedures, potentially skewing the participant pool. Moreover, a noteworthy imbalance is evident in the geographical distribution of the studies, with 23 out of 31 conducted in Asian countries. This disparity introduces the possibility of additional selection bias, as cultural and societal variations could influence the outcomes. Furthermore, a critical observation arises from the fact that none of the 31 studies systematically addressed confounding variables or adjusted data using established methods such as propensity score matching. The absence of rigorous control measures raises concerns about the reliability and validity of the findings, emphasizing the need for future research to employ more robust methodologies. Pre-clinical studies have demonstrated that pharmacological inhibition or deletion of FAP-positive CAFs results in attenuation of tumor growth and increased survival in pancreatic cancer models (14,15). However, further studies are needed to prove whether FAP overexpression has causal effects in GI cancers in general.

Our study re-evaluates and highlights the possible importance of FAP and CAFs in GI tumors. In the future, it is important to design and establish studies to evaluate the clinical significance of FAP targeting strategies with e.g., novel selective FAP inhibitors, such as FAPI-04, FAPI-46 (59), FAPI-74 (60) as well as 3BP-3940 (61), molecular imaging with 68Ga-FAPI-positron emission tomography/computed tomography (PET/CT) (59), and chimeric antigen receptor (CAR) T cells therapy (62).

Conclusions

In the current study, we showed that patients without FAP overexpression exhibited a favorable OS. Subgroup analyses revealed consistent results across esophageal, pancreatic, colorectal, and gastric cancers. This meta-analysis underscores the potential adverse prognostic significance of FAP expression in GI cancers. The variability in FAP expression and its impact on survival across different GI cancers calls for further investigation into its role and the development of selective FAP inhibitors or targeted therapies. Future research should focus on delineating the precise roles and clinical implications of FAP in GI cancers. Standardization in measuring FAP expression and controlling for confounding factors in studies are critical steps to improve the reliability of research in this area.

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Conflicts of Interest: All authors have completed the ICMJE uniform disclosure form (available at https://jgo.amegroups.com/article/view/10.21037/jgo-24-374/coif). The authors have no conflicts of interest to declare.

Ethical Statement: The authors are accountable for all aspects of the work in ensuring that questions related to the accuracy or integrity of any part of the work are appropriately investigated and resolved.

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Appendix 1

Search strategy

PubMed

"surface expressed protease"[tw] OR

- "seprase"[tw] OR
- "FAPalpha"[tw] OR
- "fibroblast activation protein-alpha"[tw] OR "FAP protein"[tw] OR
- "fibroblast-activating protein"[tw] OR
- "fibroblast proliferation factor"[tw] OR
- "fibroblast activation protein, alpha"[tw] OR
- "fibroblast activation protein"[tw] OR
- "seprase protein"[tw]

AND

"neoplasms"[Mesh]

Web of Science Core Collection (all field)

("surface expressed protease" OR "seprase" OR "FAPalpha" OR "fibroblast activation protein-alpha" OR "FAP protein" OR "fibroblast activating protein" OR "fibroblast proliferation factor" OR "fibroblast activation protein, alpha" OR "fibroblast activation protein" OR "fibroblast activation protein" OR "seprase protein") AND ("tumor" OR "neoplasm" OR "tumors" OR "neoplasia" OR "neoplasias" OR "cancer" OR "cancers" OR "malignant neoplasm" OR "neoplasm, malignant" OR "neoplasms, malignant")

Cochrane library (title, abstract, keyword)

("surface expressed protease" OR "seprase" OR "FAPalpha" OR "fibroblast activation protein-alpha" OR "FAP protein" OR "fibroblast activating protein" OR "fibroblast proliferation factor" OR "fibroblast activation protein, alpha" OR "fibroblast activation protein" OR "fibroblast activation protein" OR "seprase protein") AND ("tumor" OR "neoplasm" OR "tumors" OR "neoplasia" OR "neoplasias" OR "cancer" OR "cancers" OR "malignant neoplasm" OR "malignancy" OR "malignancies" OR "malignant neoplasms" OR "neoplasm, malignant" OR

ClinicalTrials.gov

Condition or disease:

Neoplasms

Other terms:

"surface expressed protease" OR "seprase" OR "FAPalpha" OR "fibroblast activation protein-alpha" OR "FAP protein" OR "fibroblast-activating protein" OR "fibroblast proliferation factor" OR "fibroblast activation protein"

	FAP overexpre	ession	No FAP overexp	pression		Odds Ratio	Odds Ratio		
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% Cl	M–H, Fixed, 95% Cl		
1.3.3 Esophageal car	ncer								
Ha 2014	57	71	40	45	10.1%	0.51 [0.17, 1.53]			
Higashino 2019	26	31	39	39	6.3%	0.06 [0.00, 1.15]	· · · · · · · · · · · · · · · · · · ·		
Kashima 2019	34	50	22	44	7.8%	2.13 [0.92, 4.91]			
Li 2020	37	45	63	76	8.7%	0.95 [0.36, 2.52]			
Subtotal (95% CI)		197		204	32.8%	0.93 [0.56, 1.53]	•		
Total events	154		164						
Heterogeneity: Chi ² =	= 8.22, df = 3 (P	= 0.04);	$I^2 = 63\%$						
Test for overall effect	:: Z = 0.30 (P = 0).76)							
1.3.4 Pancreatic can	cer								
Cohen 2008	18	33	9	11	6.4%	0.27 [0.05, 1.43]			
Kawase 2015	12	29	14	19	10.3%	0.25 [0.07, 0.89]			
Lo 2017	53	93	23	38	14.6%	0.86 [0.40, 1.86]			
Park 2017	71	77	65	78	5.2%	2.37 [0.85, 6.59]			
Shi 2012	36	98	25	31	25.0%	0.14 [0.05, 0.37]	_		
Wen 2019	23	33	18	34	5.6%	2.04 [0.75, 5.57]			
Subtotal (95% CI)		363		211	67.2%	0.66 [0.45, 0.97]	\bullet		
Total events	213		154						
Heterogeneity: Chi ² =	= 24.35, df = 5 (I	P = 0.00	02); I ² = 79%						
Test for overall effect	Z = 2.13 (P = 0)).03)							
		560		415	100.0%				
Total (95% CI)		500		415	100.0%	0.75 [0.55, 1.01]	-		
I otal events	367		318						
Heterogeneity: Chi ² =	= 33.93, df = 9 (F	P < 0.00	01); $I^2 = 73\%$				0.01 0.1 1 10 100		
Test for overall effect	Z = 1.88 (P = 0)).06)					Favours [experimental] Favours [control]		
Test for subgroup differences: $Chi^2 = 1.10$, $df = 1$ (P = 0.29), $I^2 = 9.2\%$									

Figure S1 Forest plot of OR with 95% CI for 1-year survival rate. The pooled ORs for esophageal and pancreatic cancer was calculated by using the random effect model. The ORs presented are FAP overexpression *vs.* no FAP overexpression. Heterogeneity was presented as I². FAP, fibroblast-activation protein; M-H, Mantel-Haenszel; CI, confidence interval; OR, odds ratio.

	FAP overexpres	sion	No FAP overexpression	on		Odds Ratio	Odds Ratio				
Study or Subgroup	Events	Total	Events T	otal	Weight	M-H, Fixed, 95% Cl	M–H, Fixed, 95% Cl				
1.4.3 Esophageal ca	ncer										
Ha 2014	41	71	36	45	13.8%	0.34 [0.14, 0.81]	.				
Higashino 2019	23	31	38	39	6.4%	0.08 [0.01, 0.64]	·				
Kashima 2019	26	50	22	44	8.3%	1.08 [0.48, 2.44]					
Li 2020	20	45	50	76	15.3%	0.42 [0.20, 0.89]					
Subtotal (95% CI)		197		204	43.7%	0.47 [0.30, 0.73]	\bullet				
Total events	110		146								
Heterogeneity: Chi ² =	= 7.49, df = 3 (P =	0.06);	$I^2 = 60\%$								
Test for overall effect	Z = 3.37 (P = 0.0)	007)									
1.4.4 Pancreatic can	cer										
Cohen 2008	10	33	7	11	5 4%	0 25 [0 06 1 04]					
Kawase 2015	3	29	7	19	5.6%	0.20 [0.04, 0.90]					
Lo 2017	19	93	, 14	38	11 7%	0.44 [0.19, 1.01]					
Park 2017	64	77	38	78	4 7%	5 18 [2 46 10 90]					
Shi 2012	7	98	24	31	25.0%	0.02 [0.01 0.07]	← ∎───				
Wen 2019	17	33	11	34	3.9%	2.22 [0.82, 5.99]					
Subtotal (95% CI)		363		211	56.3%	0.73 [0.52, 1.03]	•				
Total events	120		101								
Heterogeneity: Chi ² =	= 73.83, df = 5 (P -	< 0.000	$(001); I^2 = 93\%$								
Test for overall effect	Z = 1.79 (P = 0.0))7)									
Total (95% CI)		560		415	100.0%	0.62 [0.47, 0.81]	•				
Total events	230		247								
Heterogeneity: Chi ² =	Heterogeneity: $Chi^2 = 83.07$, $df = 9 (P < 0.0001)$; $l^2 = 89\%$										
Test for overall effect	Z = 3.52 (P = 0.0)	004)					0.01 0.1 1 10 100				
Test for subgroup dif	ferences: Chi ² = 2	.44, df	$= 1 (P = 0.12), I^2 = 59$.0%			Favours (experimental) Favours (control)				

Figure S2 Forest plot of OR with 95% CI for 2-year survival rate. The pooled ORs for esophageal and pancreatic cancer was calculated by using the random effect model. The ORs presented are FAP overexpression *vs.* no FAP overexpression. Heterogeneity was presented as I². FAP, fibroblast-activation protein; M-H, Mantel-Haenszel; CI, confidence interval; OR, odds ratio.

	FAP overexpre	ession	No FAP overexp	ression		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% Cl	M–H, Fixed, 95% Cl
1.5.3 Esophageal can	cer						
Ha 2014	37	71	31	45	11.8%	0.49 [0.22, 1.08]	
Higashino 2019	20	31	35	39	7.1%	0.21 [0.06, 0.74]	
Kashima 2019	25	50	18	44	6.2%	1.44 [0.64, 3.27]	
Li 2020	12	45	43	76	15.2%	0.28 [0.13, 0.62]	
Subtotal (95% CI)		197		204	40.3%	0.51 [0.33, 0.77]	\bullet
Total events	94		127				
Heterogeneity: Chi ² =	10.33, df = 3 (F	P = 0.02)	; $I^2 = 71\%$				
Test for overall effect:	Z = 3.17 (P = 0)	.002)					
1.5.4 Pancreatic canc	er						
Cohen 2008	4	33	7	11	6.0%	0.08 [0.02, 0.40]	
Kawase 2015	0	29	4	19	3.4%	0.06 [0.00, 1.16]	• • • • • • • • • • • • • • • • • • • •
Lo 2017	10	93	9	38	7.4%	0.39 [0.14, 1.05]	
Park 2017	7	77	64	78	37.5%	0.02 [0.01, 0.06]	←∎──
Wen 2019	0	33	8	34	5.4%	0.05 [0.00, 0.84]	<
Subtotal (95% CI)		265		180	59.7%	0.08 [0.04, 0.14]	◆
Total events	21		92				
Heterogeneity: Chi ² =	16.79, df = 4 (F)	P = 0.002	2); I ² = 76%				
Test for overall effect:	Z = 8.99 (P < 0)	.00001)					
Total (95% CI)		462		384	100.0%	0 25 [0 18 0 35]	
Total overts	115	102	210	504	100.0/0	0.25 [0.10, 0.55]	•
Hotorogonoity: Chi ² -	115 10 00 df _ 0 (r	~ ~ ^ ^ ^	219				
Test for everall effects	49.09, 01 = 8 (P < 0)	< 0.000	501), 1 = 84%				0.01 0.1 1 10 100
Test for subgroup diff	L = 0.40 (P < 0)	27.04 -	If _ 1 (D < 0.0000	1) $1^2 - 0^2$. 10/		Favours [experimental] Favours [control]

Figure S3 Forest plot of OR with 95% CI for 3-year survival rate. The pooled ORs for esophageal and pancreatic cancer was calculated by using the random effect model. The ORs presented are FAP overexpression *vs.* no FAP overexpression. Heterogeneity was presented as I². FAP, fibroblast-activation protein; M-H, Mantel-Haenszel; CI, confidence interval; OR, odds ratio.

	FAP overexpre	ession	No FAP overexp	ression		Odds Ratio	Odds Ratio		
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% Cl	M–H, Fixed, 95% Cl		
1.6.3 Esophageal ca	ncer								
Ha 2014	29	71	28	45	32.0%	0.42 [0.19, 0.90]	_		
Higashino 2019	16	31	34	39	23.0%	0.16 [0.05, 0.51]			
Kashima 2019	22	50	17	44	16.0%	1.25 [0.55, 2.85]			
Li 2020	34	45	11	76	3.2%	18.26 [7.18, 46.43]			
Subtotal (95% CI)		197		204	74.0%	1.28 [0.87, 1.87]	•		
Total events	101		90						
Heterogeneity: Chi ² =	= 51.62, df = 3 (F	o < 0.00	001 ; $I^2 = 94\%$						
Test for overall effect	t: $Z = 1.26 (P = 0)$).21)							
1.6.4 Pancreatic can	cer								
Cohen 2008	1	33	7	11	16.0%	0.02 [0.00, 0.19]	← ∎		
Kawase 2015	0	29	0	19		Not estimable			
Lo 2017	2	93	2	38	4.4%	0.40 [0.05, 2.92]			
Park 2017	64	77	21	78	5.6%	13.36 [6.14, 29.10]			
Subtotal (95% CI)		232		146	26.0%	2.93 [1.69, 5.09]	•		
Total events	67		30						
Heterogeneity: Chi ² =	= 36.72, df = 2 (F	o < 0.00	001 ; $I^2 = 95\%$						
Test for overall effect	t: $Z = 3.82 (P = 0)$.0001)							
Total (95% CI)		429		350	100.0%	1.71 [1.26, 2.32]	\bullet		
Total events	168		120						
Heterogeneity: $Chi^2 = 97.64$, $df = 6$ (P < 0.00001); $l^2 = 94\%$									
Test for overall effect	t: Z = 3.42 (P = 0)	.0006)					U.UI U.I I IO 100		
Test for subgroup differences: $Chi^2 = 5.93$, $df = 1$ (P = 0.01), $l^2 = 83.1\%$									

Figure S4 Forest plot of OR with 95% CI for 5-year survival rate. The pooled ORs for esophageal and pancreatic cancer was calculated by using the random effect model. The ORs presented are FAP overexpression *vs.* no FAP overexpression. Heterogeneity was presented as I². FAP, fibroblast-activation protein; M-H, Mantel-Haenszel; CI, confidence interval; OR, odds ratio.